The Effect of ESG on the Global Equities Lending Market

Institutional ownership, securities lending supply, borrowing demand, and investor engagement are significantly impacted by ESG preferences

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It is well known that ESG has become an important factor in investment decision-making, but to what degree has it directly influenced the holdings of institutional investors and their engagement over proxy votes? What about the borrowing demand of short sellers?

The paper looks to answer these questions with a unique empirical study of ESG's effect on the global equities lending market.



By combining equities lending, ESG, and proxy vote data, we quantify the impact ESG has had on lending supply, short selling demand, and institutional investor engagement.

Our findings suggest that ESGconsiderations are deeply embedded in the securities lending market and are growing in importance.

Importantly, the analysis identifies ways in which institutional investors can balance the revenue earned through securities lending with the pursuit of their ESG objectives.

Key Findings



ESG significantly impacts securities lending supply

A significant positive relationship exists between a company's performance on ESG attributes and the shares available on the securities lending market.



Evidence of "green shorting"

Firms that perform poorly on material ESG attributes are associated with significantly higher fees and increased levels of borrowing demand, a proxy for short selling.



ESG influences institutional engagement

Trends in securities lending share recalls over proxy record dates suggest institutions engage more with companies with poor ESG characteristics.



Little evidence of "empty voting"

A concern cited by some beneficial owners; we find minimal evidence of borrowing securities to increase short-term influence over proxy votes.



Sourcing ESG, Securities Lending, and Institutional Ownership Data

The underlying data used in our study is sourced from several data sets. Firm-level ESG scores are produced by MSCI and measure a company's resilience to long-term industry material ESG risks. Higher scores indicate better performance in mitigating ESG risks material to that industry. Given there are many ways in which we can measure ESG, we also use ESG sentiment scores sourced from TrueValueLabs for further testing.

We use IHS Markit® to source daily global equities lending data, including fees, supply, demand, and utilization. Institutional ownership is captured by Thomson-Reuters Institutional Holdings (13F) Database. The combined study covers roughly 3,500 global equities from 2015 to 2021, including more than 33,000 proxy record dates.

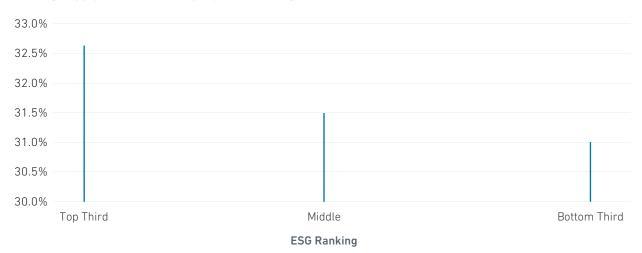


ESG Performance of a Firm Directly Impacts Securities Lending Supply

To start our analysis, we examine how a firm's performance on ESG attributes impacts the average number of shares available in the lending market. In other words, is the ESG performance of a firm correlated with its level of lending supply? Taking a high-level approach, we rank the yearly average ESG scores of each company and group them into terciles - the bottom tercile being the poorest performing. It is clear in **Figure 1** that companies performing poorly on material ESG attributes have fewer shares available for lending relative to their market cap – close to 1.5% on average.

Figure 1: ESG is a significant driver of securities lending supply

Lending Supply (% of Market Cap) by ESG Ranking



Source: State Street Global Markets, MSCI, IHS Markit, Thomson Reuters. 2015 - 2021

¹ MSCI ESG Fund Ratings Summary (2021). MSCI ESG Research LLC.

As many readers will note, a number of market factors explain differences in securities lending supply. For example, it is well understood that market cap, concentration of institutional ownership, and other factors can contribute to supply differences. To understand if a company's ESG performance explains differences in securities lending supply above and beyond that of typical equity factors, we run a panel regression controlling for year and firm fixed effects. We include a number of common controls, such as market cap. institutional ownership concentration, book-tomarket, and momentum. We find a statistically significant positive relationship between ESG performance and securities lending supply. The results can be found in **Appendix**.

What explains this relationship? Digging deeper, we find two contributing factors. The biggest contributor, not surprisingly, is the level of

institutional ownership. It is well understood that institutional ownership is a significant predictor of lending supply as institutional investors are the primary participants in lending programs. As ESG has become an important factor of investment decisions, institutions have increasingly shifted ownership away from stocks that perform poorly on ESG characteristics to stocks that are considered more sustainable. This trend results in Figure 2, in which we observe the average difference in institutional ownership relative to market cap for top ESG performing firms versus the bottom performers each year.

In 2021, a firm in the top third of ESG scores had, on average, 8% more of its market cap owned by institutional investors relative to a firm in the bottom third. This is 16-times greater than the equivalent measure in 2015, indicating ESG is of growing importance in explaining institutional holdings.

Figure 2: Institutional investors are paying more attention to ESG performance of firms

Difference in Institutional Ownership (Top Minus Bottom ESG Companies)



Source: State Street Global Markets, MSCI, IHS Markit, Thomson Reuters.

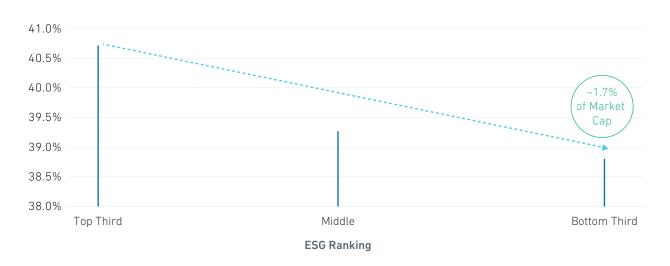
The preference of institutional investors to hold positively ranked ESG stocks is not the only driver of differences in securities lending supply. Investors seem to also restrict the lending of shares of poorly performing ESG companies. This is apparent when we observe the level of securities lending supply relative to institutional ownership in **Figure 3**.

Not only do institutional investors own fewer shares of poorly performing ESG firms but we also find they are less willing to lend those shares out. The average percent of institutional ownership that is available on the securities lending market for the poorly ranked (bottom third) ESG firms is 39%, relative to 40% for the top third.

Supply is only one side of the securities lending market. For a full picture of how ESG effects investor preferences, we turn our analysis to borrower demand.

Figure 3: Institutional investors restrict lending on holdings of stocks with poor ESG performance

Supply as a Percent of Institutional Ownership



 ${\tt Source: State \ Street \ Global \ Markets, \ MSCI, \ IHS \ Markit, \ Thomson \ Reuters.}$



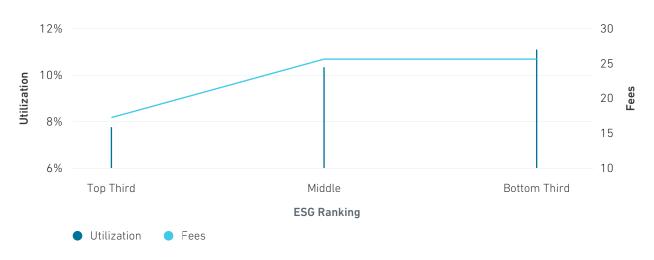
ESG Effects on Short Selling Demand

The positive relationship between ESG rankings and securities lending supply raises interesting questions. For example, given lower supply levels for poorly ranked ESG stocks, are these stocks more expensive to short? What about short sellers, do they demand more of these stocks? Does this mean lenders can earn more on them?

Let's start with what we expect to find. With lower levels of aggregate supply and demand being equal, we expect poorly ranked ESG stocks to be more expensive to short, thus providing an opportunity for lenders to earn higher revenue.

Figure 4: Higher utilization and fees for poorly ranked ESG firms

Fees and Utilization by ESG Ranking



Source: State Street Global Markets, MSCI, IHS Markit, Thomson Reuters. 2015 - 2021

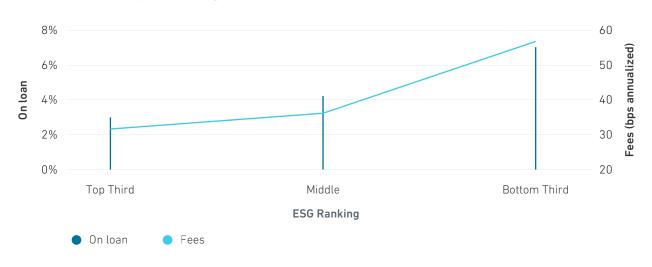
Indeed, this is what we observe in **Figure 4**. Utilization — the ratio of shares on loan versus total supply — for poorly-performing ESG firms is significantly greater than that of good performers. This in turn impacts borrowing fees, which are also almost twice as high for the bottom third of poorly ranked ESG firms.

As with supply, the difference in demand between positively and negatively ranked ESG firms increased substantially over the years since 2015.

However, it is unclear if the results in **Figure 4** are driven by differing levels in supply or if it is also due to differences in demand.

Figure 5: Energy sector, short sellers identify firms performing poorly on ESG attributes

Fees and Utilization by ESG Ranking



Source: State Street Global Markets, MSCI, IHS Markit, Thomson Reuters.

2015 - 2021

In **Figure 5**, we isolate demand as a proportion of market cap. Specifically, we view this through the energy sector. Our evidence suggests that there is considerable effort exerted in determining unsustainable firms within ESG risky industries or sectors. For example, within the energy sector, companies in the bottom third of ESG scores are nearly twice as expensive and have double the level of shorting demand compared to their peers; suggesting that green shorting is an active endeavor rather than blanket selling of ESG-risky industries.

As with supply, there are a number of factors that contribute to differences in shorting demand (onloan value). Industry hedging demand, market cap, short-term momentum are example market factors that can explain differences in borrowing demand. Controlling for these factors, we find that ESG performance explains differences in borrowing demand above and beyond documented market factors. The results are presented in the **Appendix**.



There are Significant Changes in Supply and Demand Dynamics Over Proxy Record Dates

A central tenant of responsible asset stewardship is engagement with shareholder investments. Having uncovered how a firm's performance on ESG attributes influences aggregate securities lending, we want to understand how the supply level changes surrounding proxy votes.

Proxy events serve as an important channel for institutions to exert their influence and long-term company vision, a process of which is often required by law and scrutinized by third parties. This is also a core principle of responsible asset stewardship.

To participate in a proxy vote, firms must hold the shares of a company on the date of proxy record, called the proxy record date. As a result, it is critical to understand the market dynamics of securities lending leading up to and after the proxy record dates.

Institutional investors wanting to engage in a proxy vote must recall or restrict their supply of shares **before the proxy record date**. As we expect, we observe a significant decrease in available lending supply within 30 days leading up to the record date. Immediately after the record date, supply snaps back, accounting for a 1% increase in lending supply relative to a firm's market cap, on average.

Figure 6: Supply dynamics around proxy record dates

28.0% 12.2% 27.8% 27.5% 12.0% 27.3% 11.8% 27.0% 26.8% 11.6% 26.5% 11.4% 26.3% 26.0% 11.2% 30 26 22 26 30

Days Relative to Record Date

Source: State Street Global Markets, MSCI, IHS Markit, Thomson Reuters.

Utilization (rhs)

2015 - 2021

Supply

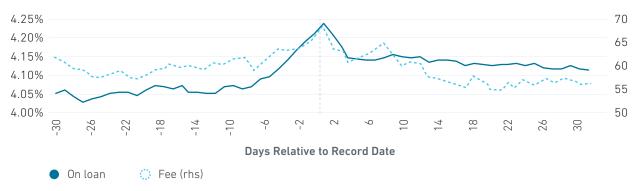
Record Date (T0)

On the other hand, demand is observed to rise approaching the date of record, as illustrated in **Figure 7**. At face value, this spurs a concern of malicious "empty voting" practices, in which an

investor can increase their short-term influence on a vote by borrowing shares over the proxy record date. This would be a concern for asset owners with a long-term mindset.

Figure 7: Demand dynamics around record dates

Record Date (T0)



Source: State Street Global Markets, MSCI, IHS Markit, Thomson Reuters. 2015 - 2021

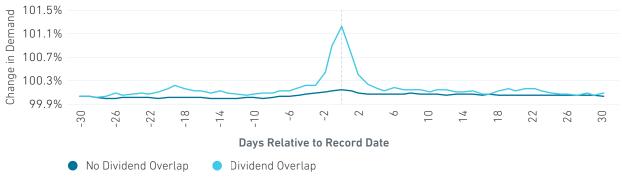
However, as we dig deeper, we find evidence rejecting the notion that increased demand is driven by empty voting practices. In **Figure 8** we breakout record dates into two groups:

(1) those with overlapping ex-dividend dates and (2) those that only have a proxy record date.

Increases in demand are negligible for proxy-only record dates (i.e., driven by ex-dividend dates), alleviating concerns that lending facilitates the expression of views not aligned with longer-term asset stewardship endeavors.

Figure 8: No evidence of 'empty voting'

Record Date (T0)

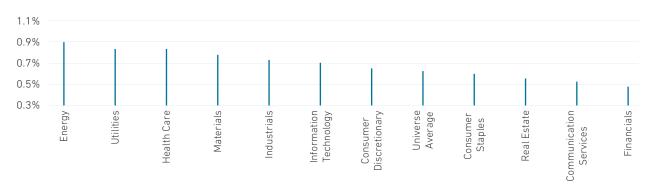


Source: State Street Global Markets, MSCI, IHS Markit, Thomson Reuters.

Next, we examine how ESG might influence the level of institutional engagement. At the sector level (GICS level 1), we expect to observe increased engagement (proxied by changes in supply leading up to proxy record date) in sectors that are "at risk" of negative ESG-stocks. Indeed, we find that sectors such as Energy, Utilities, and Materials have twice the level of engagement than those less at risk.

Figure 9: Average firm engagement by sector

Institutional Engagement Over Proxy Record Dates by Sector



Source: State Street Global Markets, MSCI, IHS Markit, Thomson Reuters.

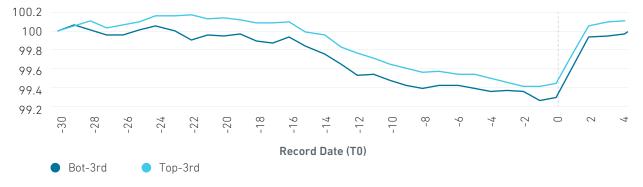
2015 - 2021

We also expect the average level of institutional engagement to be higher for firms performing poorly on their ESG attributes, as we expect investors to push for more positive ESG

outcomes through their voting power. **Figure 10** presents the average decreases in supply (level of recall/ restrictions) for the bottom third and top third of ESG ranking firms.

Figure 10: ESG-driven recall

ESG and Supply Restrictions



Source: State Street Global Markets, MSCI, IHS Markit, Thomson Reuters.

At first glance, it seems like institutions may not be considering the ESG performance of a firm when engaging, as there was only a marginal difference in supply recalls/ restrictions for the poorly ranked firms. However, it is only part of the story.

While it is true that there is only a marginal increase in engagement on firms with poor ESG rankings, when we consider the lending revenue that can be earned over those periods, a different story appears. Poorly ranked ESG firms earn twice the level of revenue over proxy record dates, Figure 11. This suggests institutional investors are willing to forgo twice the amount of revenue in order to engage with those firms.

Figure 11: Opportunity cost of engagement

Lending Revenue over Proxy Record Dates (annualized)



 ${\tt Source: State \ Street \ Global \ Markets, \ MSCI, \ IHS \ Markit, \ Thomson \ Reuters.}$



Conclusion



We find evidence that ESG considerations play an important role in institutional investment decisions, which in turn, has significant implications on the securities lending market. There are lower levels of institutional ownership, increased levels of shorting, and more engagement for firms that perform poorly on material ESG characteristics. We observe these trends strengthening through time as institutional investors appear to be place greater emphasis on ESG characteristics.

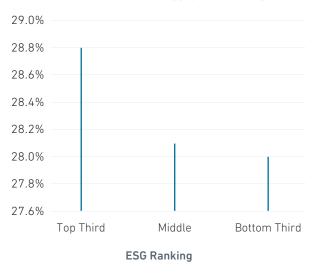
Our analysis provides insights into how asset owners and managers are balancing the incremental revenue earned through securities lending with their ESG-objectives. We hope that this helps market participants think through their securities lending program and understand how ESG is impacting the market as a whole.

Appendix

Supply Trends in the Small Cap Universe

We find similar trends in securities lending supply with the small cap universe.

ESG Scores and Lendable Supply - Small Caps



Source: State Street Global Markets, MSCI, IHS Markit, Thomson Reuters.

2015 - 2021

Regression Controls

To test for statistical significance we run multivariate ordinary least squares (OLS) regressions, including a constant and our control variables. We control for equity factors adding in the log of market capitalization (SIZE), bookto-market ratio (BM), turnover (TURNOVER), bid-ask spread (SPREAD), short-term momentum (SHORT-TERM MOM) measured as a stock's

cumulative return over the last 5 trading days, and long-term momentum (LONG-TERM MOM), the cumulative return over the previous 252 trading days. We add dummy variables to control for penny stocks worth less than five dollars (PRICE < \$5) and signify the proxy record date (RDATE). We account for institutional ownership (INST), and ownership concentration (INST CONC) measured by Herfindahl index. All our regressions contain year-fixed and firm-fixed effects to each of regression, and include robust standard errors clustered at firm level. Year-fixed effects are accomplished through the inclusion of year dummy variables (i.e., 2016_DUMMY).

Supply Regression

Using SUPPLY as the dependent variable along with our set of control variables, our results found that the proxy record date (RDATE) was a statistically significant variable in describing the amount of lendable supply. All else equal, proxy record dates result in a ~1.23% decrease in the lending supply for firms.

Adding ESG to our regression, all else equal, we expect a firm with an ESG score that is 1 point higher, to have ~.3% more of its market cap available for lending.

As with all our regressions, we implemented firm-fixed and year-fixed effects to ensure robust standard errors.

Explaining Supply With Market Factors

| OLS Regression Results | | | | |
|------------------------|---------|---------|----------|-------|
| SUPPLY | | | | |
| | COEF | STD ERR | t | P> t |
| CONSTANT | 0.1606 | 0.040 | 4.017* | 0.000 |
| RDATE | -0.0123 | 0.001 | -13.911* | 0.000 |
| INST | 0.2858 | 0.010 | 29.130* | 0.000 |
| INST CONC | 0.0849 | 0.009 | 9.006* | 0.000 |
| SIZE | -0.0266 | 0.010 | -2.714* | 0.007 |
| ВМ | 0.0110 | 0.003 | 3.751* | 0.000 |
| TURNOVER | -0.2021 | 0.148 | -1.369 | 0.171 |
| SPREAD | 0.0000 | 0.000 | -0.577 | 0.564 |
| PRICE<\$5 | -0.0657 | 0.003 | -7.772* | 0.000 |
| Short-Term Mom | -0.0343 | 0.003 | -12.759* | 0.000 |
| Long-Term Mom | -0.0005 | 0.001 | -0.372 | 0.710 |
| 2016_dummy | 0.0087 | 0.002 | 5.703* | 0.000 |
| 2017_dummy | 0.0239 | 0.002 | 10.674* | 0.000 |
| 2018_dummy | 0.0608 | 0.003 | 20.994* | 0.000 |
| 2019_dummy | 0.0694 | 0.003 | 26.628* | 0.000 |
| 2020_dummy | 0.0998 | 0.003 | 37.037* | 0.000 |
| 2021_dummy | 0.1256 | 0.003 | 43.413* | 0.000 |
| Firm FE | Yes | | | |
| Year FE | Yes | | | |
| Adj. R-Squared | 0.525 | | | |
| Number of Firms | 3,469 | | | |

 $\mathsf{SUPPLY}_{\mathsf{t}} = \alpha + \beta_{\mathsf{t}} \mathsf{Record} \; \mathsf{Date} + \delta \mathsf{Controls}_{\mathsf{t}} + \epsilon$

Including ESG Metrics

| OLS Regression Results | | | | |
|------------------------|---------|---------|----------|--------|
| | | SUPPLY | | |
| | COEF | STD ERR | t | P> t |
| CONSTANT | 0.1594 | 0.040 | 3.976* | 0.000 |
| RDATE | -0.0121 | 0.001 | -13.863* | 0.000 |
| INST | 0.2893 | 0.010 | 29.630* | 0.000 |
| INST CONC | 0.0819 | 0.009 | 8.715* | 0.000 |
| SIZE | -0.0300 | 0.010 | -3.047* | 0.002 |
| ВМ | 0.0117 | 0.003 | 3.969* | 0.000 |
| TURNOVER | -0.1943 | 0.143 | -1.358 | -1.358 |
| SPREAD | 0.0000 | 0.000 | -0.534 | 0.593 |
| PRICE<\$5 | -0.0635 | 0.008 | -7.533* | 0.000 |
| Short-Term Mom | -0.0341 | 0.003 | -12.790* | 0.000 |
| Long-Term Mom | -0.0004 | 0.001 | -0.319 | 0.750 |
| MSCI ESG SCORE | 0.0030 | 0.001 | 3.601* | 0.000 |
| 2016_dummy | 0.0085 | 0.002 | 5.611* | 0.000 |
| 2017_dummy | 0.0236 | 0.002 | 10.585* | 0.000 |
| 2018_dummy | 0.0598 | 0.003 | 20.420* | 0.000 |
| 2019_dummy | 0.0682 | 0.003 | 25.778* | 0.000 |
| 2020_dummy | 0.0983 | 0.003 | 35.681* | 0.000 |
| 2021_dummy | 0.1239 | 0.003 | 41.763* | 0.000 |
| Firm FE | Yes | | | |
| Year FE | Yes | | | |
| Adj. R-Squared | 0.528 | | | |
| Number of Firms | 3,469 | | | |

 $\text{SUPPLY}_{_{t}} = \alpha + \beta_{_{1}} \text{MSCI ESG Score} + 8 \text{ Controls}_{_{t}} + \epsilon$

Demand Regression

Using ONLOAN as the dependent variable along with our set of control variables, our results found that the proxy record date (RDATE) was a statistically significant variable in describing the amount of demanded supply. All else equal, proxy record dates result in a ~0.14% increase in demand for shares.

This raised concerns that there may be evidence of the practice of "Empty Voting", a major criticism of securities lending practices. To investigate, we control for ex-dividend date through the addition of a dividend date dummy

variable (DDATE), in our ONLOAN regression. With the inclusion of the dividend date dummy (DDATE), the proxy record date (RDATE) becomes insignificant in explaining demand increases. All else equal, dividend record dates result in a ~0.41% increase in demand for shares.

Adding ESG to our regression, all else equal, we expect a firm with an ESG score that is 1 point lower, to have ~.4% more of its market cap shorted.

As with all our regressions, we implemented firm-fixed and year-fixed effects to ensure robust standard errors.

increase in demand for shares as per proxy record date.

~0.14% ~0.41%

increase in demand for shares as per dividend record date.

Explaining Demand With Market Factors

| OLS Regression Results | | | | |
|------------------------|---------|---------|----------|-------|
| | | ONLOAN | | |
| | COEF | STD ERR | t | P> t |
| CONSTANT | 0.2935 | 0.021 | 14.087* | 0.000 |
| RDATE | 0.0014 | 0.000 | 2.865* | 0.004 |
| INST | 0.0245 | 0.004 | 6.463* | 0.000 |
| INST CONC | -0.0061 | 0.004 | -1.713 | 0.087 |
| SIZE | -0.0620 | 0.005 | -13.324* | 0.000 |
| ВМ | 0.0024 | 0.002 | 1.517 | 0.129 |
| TURNOVER | 0.1906 | 0.136 | 1.403 | 0.161 |
| SPREAD | 0.0000 | 0.000 | -6.382* | 0.000 |
| PRICE<\$5 | 0.0157 | 0.004 | 4.208* | 0.000 |
| Short-Term Mom | -0.0101 | 0.002 | -5.469* | 0.000 |
| Long-Term Mom | -0.0016 | 0.001 | -1.723 | 0.085 |
| 2016_dummy | 0.0024 | 0.001 | 2.337* | 0.020 |
| 2017_dummy | -0.0013 | 0.001 | -0.921 | 0.357 |
| 2018_dummy | -0.0018 | 0.002 | -1.037 | 0.300 |
| 2019_dummy | 0.0001 | 0.002 | 0.084 | 0.933 |
| 2020_dummy | -0.0036 | 0.001 | -2.474* | 0.013 |
| 2021_dummy | -0.0061 | 0.002 | -4.029* | 0.000 |
| Firm FE | Yes | | | |
| Year FE | Yes | | | |
| Adj. R-Squared | 0.115 | | | |
| Number of Firms | 3,469 | | | |

 $\mathbf{ONLOAN} = \alpha + \beta_1 \mathbf{Record\ Date} + \delta \mathbf{Controls}_{_{t}} + \epsilon$

Demand Regression With Ex-Dividend Date

| OLS Regression Results | | | | |
|------------------------|---------|---------|----------|-------|
| | | ONLOAN | | |
| | COEF | STD ERR | t | P> t |
| CONSTANT | 0.2936 | 0.019 | 15.275* | 0.000 |
| RDATE | 0.0007 | 0.000 | 1.447 | 0.148 |
| DDATE | 0.0041 | 0.001 | 3.087* | 0.002 |
| INST | 0.0583 | 0.004 | 14.134* | 0.000 |
| INST CONC | 0.0126 | 0.004 | 3.494* | 0.000 |
| SIZE | -0.0664 | 0.004 | -15.158* | 0.000 |
| ВМ | 0.0013 | 0.001 | 1.105 | 0.269 |
| TURNOVER | 0.1393 | 0.056 | 2.467* | 0.014 |
| SPREAD | 0.0000 | 0.000 | -7.737* | 0.000 |
| PRICE<\$5 | 0.0176 | 0.003 | 6.192* | 0.000 |
| Short-Term Mom | -0.0115 | 0.002 | -5.317* | 0.000 |
| Long-Term Mom | -0.0016 | 0.001 | -2.454* | 0.014 |
| 2016_dummy | 0.0031 | 0.001 | 3.052* | 0.002 |
| 2017_dummy | 0.0003 | 0.001 | 0.228 | 0.819 |
| 2018_dummy | 0.0000 | 0.002 | -0.025 | 0.900 |
| 2019_dummy | 0.0013 | 0.002 | 0.791 | 0.429 |
| 2020_dummy | -0.0042 | 0.001 | -2.837* | 0.005 |
| 2021_dummy | -0.0084 | 0.002 | -5.524* | 0.000 |
| Firm FE | Yes | | | |
| Year FE | Yes | | | |
| Adj. R-Squared | 0.145 | | | |
| Number of Firms | 3,469 | | | |

 $\mathbf{ONLOAN_t} = \alpha + \beta_1 \mathbf{Record} \ \mathbf{Date} + \beta_2 \mathbf{Ex} \ \mathbf{Div.} \ \mathbf{Date} + 8 \mathbf{Controls_t} + \epsilon$

Demand Regression With ESG

| OLS Regression Results | | | | |
|------------------------|---------|---------|----------|-------|
| ONLOAN | | | | |
| | COEF | STD ERR | t | P> t |
| CONSTANT | 0.3042 | 0.020 | 14.876* | 0.000 |
| RDATE | 0.0013 | 0.000 | 2.613* | 0.009 |
| INST | 0.0234 | 0.004 | 6.214* | 0.000 |
| INST CONC | -0.0043 | 0.004 | -1.203 | 0.229 |
| SIZE | -0.0603 | 0.005 | -13.101* | 0.000 |
| ВМ | 0.0020 | 0.002 | 1.280 | 0.201 |
| TURNOVER | 0.1857 | 0.133 | 1.398 | 0.162 |
| SPREAD | 0.0000 | 0.000 | -6.248* | 0.000 |
| PRICE<\$5 | 0.0142 | 0.004 | 3.849* | 0.000 |
| Short-Term Mom | -0.0102 | 0.002 | -5.559* | 0.000 |
| Long-Term Mom | -0.0017 | 0.001 | -1.760 | 0.078 |
| MSCI ESG SCORE | -0.0038 | 0.001 | -4.234* | 0.000 |
| 2016_dummy | 0.0026 | 0.001 | 2.508* | 0.012 |
| 2017_dummy | -0.0012 | 0.001 | -0.801 | 0.423 |
| 2018_dummy | -0.0008 | 0.002 | -0.473 | 0.636 |
| 2019_dummy | 0.0011 | 0.002 | 0.689 | 0.491 |
| 2020_dummy | -0.0025 | 0.001 | -1.678 | 0.093 |
| 2021_dummy | -0.0051 | 0.002 | -3.316* | 0.001 |
| Firm FE | Yes | | | |
| Year FE | Yes | | | |
| Adj. R-Squared | 0.120 | | | |
| Number of Firms | 3,469 | | | |

 $\mathbf{ONLOAN} = \alpha + \beta_1 \mathbf{MSCI} \; \mathbf{ESG} \; \mathbf{Score} + 8 \mathbf{Controls}_{t} + \epsilon$

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